

Arithmetic properties of multiplicative integer-valued perturbed random walks

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Abstract Let $(\xi_1, \eta_1), (\xi_2, \eta_2), \dots$ be independent identically distributed \mathbb{N}^2 -valued random vectors with arbitrarily dependent components. The sequence $(\Theta_k)_{k \in \mathbb{N}}$ defined by $\Theta_k = \Pi_{k-1} \cdot \eta_k$, where $\Pi_0 = 1$ and $\Pi_k = \xi_1 \cdot \dots \cdot \xi_k$ for $k \in \mathbb{N}$, is called a multiplicative perturbed random walk. Arithmetic properties of the random sets $\{\Pi_1, \Pi_2, \dots, \Pi_k\} \subset \mathbb{N}$ and $\{\Theta_1, \Theta_2, \dots, \Theta_k\} \subset \mathbb{N}$, $k \in \mathbb{N}$, are studied. In particular, distributional limit theorems for their prime counts and for the least common multiple are derived.

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1 Introduction

Let $(\xi_1, \eta_1), (\xi_2, \eta_2), \dots$ be independent copies of an \mathbb{N}^2 -valued random vector (ξ, η) with arbitrarily dependent components. Denote by $(\Pi_k)_{k \in \mathbb{N}_0}$ (as usual, $\mathbb{N}_0 := \mathbb{N} \cup \{0\}$) the standard multiplicative random walk defined by

$$\Pi_0 := 1, \quad \Pi_k = \xi_1 \cdot \xi_2 \cdots \xi_k, \quad k \in \mathbb{N}.$$

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A *multiplicative perturbed random walk* is the sequence $(\Theta_k)_{k \in \mathbb{N}}$ given by

$$\Theta_k := \Pi_{k-1} \cdot \eta_k, \quad k \in \mathbb{N}.$$

Note that if $\mathbb{P}\{\eta = \xi\} = 1$, then $\Pi_k = \Theta_k$ for all $k \in \mathbb{N}$. If $\mathbb{P}\{\xi = 1\} = 1$, then $(\Theta_k)_{k \in \mathbb{N}}$ is just a sequence of independent copies of a random variable η . In this article we investigate some arithmetic properties of the random sets $(\Pi_k)_{k \in \mathbb{N}}$ and $(\Theta_k)_{k \in \mathbb{N}}$.

To set the scene, we introduce first some necessary notation. Let \mathcal{P} denote the set of prime numbers. For an integer $n \in \mathbb{N}$ and $p \in \mathcal{P}$, let $\lambda_p(n)$ denote the multiplicity of prime p in the prime decomposition of n , that is,

$$n = \prod_{p \in \mathcal{P}} p^{\lambda_p(n)}.$$

For every $p \in \mathcal{P}$, the function $\lambda_p : \mathbb{N} \mapsto \mathbb{N}_0$ is totally additive in the sense that

$$\lambda_p(mn) = \lambda_p(m) + \lambda_p(n), \quad p \in \mathcal{P}, \quad m, n \in \mathbb{N}.$$

The set of functions $(\lambda_p)_{p \in \mathcal{P}}$ is a basic brick from which many other arithmetic functions can be constructed. For example, with $\text{GCD}(A)$ and $\text{LCM}(A)$ denoting the greatest common divisor and the least common multiple of a set $A \subset \mathbb{N}$, respectively, we have

$$\text{GCD}(A) = \prod_{p \in \mathcal{P}} p^{\min_{n \in A} \lambda_p(n)} \quad \text{and} \quad \text{LCM}(A) = \prod_{p \in \mathcal{P}} p^{\max_{n \in A} \lambda_p(n)}.$$

The listed arithmetic functions applied either to $A = \{\Pi_1, \dots, \Pi_n\}$ or $A = \{\Theta_1, \dots, \Theta_n\}$ are the main objects of investigation in the present paper. From the additivity of λ_p we infer

$$S_k(p) := \lambda_p(\Pi_k) = \sum_{j=1}^k \lambda_p(\xi_j), \quad p \in \mathcal{P}, \quad k \in \mathbb{N}_0, \tag{1}$$

and

$$T_k(p) := \lambda_p(\Theta_k) = \sum_{j=1}^{k-1} \lambda_p(\xi_j) + \lambda_p(\eta_k), \quad p \in \mathcal{P}, \quad k \in \mathbb{N}. \tag{2}$$

Fix any $p \in \mathcal{P}$. Formulae (1) and (2) demonstrate that $S(p) := (S_k(p))_{k \in \mathbb{N}_0}$ is a standard additive random walk with the generic step $\lambda_p(\xi)$, whereas the sequence $T(p) := (T_k(p))_{k \in \mathbb{N}}$ is a particular instance of an *additive perturbed random walk*, see [6], generated by the pair $(\lambda_p(\xi), \lambda_p(\eta))$.

2 Main results

2.1 Distributional properties of the prime counts $(\lambda_p(\xi), \lambda_p(\eta))$

As is suggested by (1) and (2) the first step in the analysis of $S(p)$ and $T(p)$ should be the derivation of the joint distribution $(\lambda_p(\xi), \lambda_p(\eta))_{p \in \mathcal{P}}$. The next lemma confirms

that the finite-dimensional distributions of the infinite vector $(\lambda_p(\xi), \lambda_p(\eta))_{p \in \mathcal{P}}$, are expressible via the probability mass function of (ξ, η) . However, the obtained formulae are not easy to handle except some special cases. For $i, j \in \mathbb{N}$, put

$$u_i := \mathbb{P}\{\xi = i\}, \quad v_j := \mathbb{P}\{\eta = j\}, \quad w_{i,j} := \mathbb{P}\{\xi = i, \eta = j\}.$$

Lemma 1. Fix $p \in \mathcal{P}$ and nonnegative integers $(k_q)_{q \in \mathcal{P}, q \leq p}$ and $(\ell_q)_{q \in \mathcal{P}, q \leq p}$. Then

$$\mathbb{P}\{\lambda_q(\xi) \geq k_q, \lambda_q(\eta) \geq \ell_q, q \in \mathcal{P}, q \leq p\} = \sum_{i,j=1}^{\infty} w_{Ki,Lj},$$

where $K := \prod_{q \leq p, q \in \mathcal{P}} q^{k_q}$ and $L := \prod_{q \leq p, q \in \mathcal{P}} q^{\ell_q}$.

Proof. This follows from

$$\begin{aligned} & \mathbb{P}\{\lambda_q(\xi) \geq k_q, \lambda_q(\eta) \geq \ell_q, q \in \mathcal{P}, q \leq p\} \\ &= \mathbb{P}\left\{ \prod_{q \leq p, q \in \mathcal{P}} q^{k_q} \text{ divides } \xi, \prod_{q \leq p, q \in \mathcal{P}} q^{\ell_q} \text{ divides } \eta \right\} = \sum_{i,j=1}^{\infty} w_{Ki,Lj}. \end{aligned}$$

Obviously, if ξ and η are independent, then

$$\sum_{i,j=1}^{\infty} w_{Ki,Lj} = \left(\sum_{i=1}^{\infty} u_{Ki} \right) \left(\sum_{j=1}^{\infty} v_{Lj} \right). \quad \square$$

We proceed with the series of examples.

Example 1. For $\alpha > 1$, let $\mathbb{P}\{\xi = k\} = (\zeta(\alpha))^{-1} k^{-\alpha}$, $k \in \mathbb{N}$, where ζ is the Riemann zeta-function. For $k \in \mathbb{N}$, $p_1, \dots, p_k \in \mathcal{P}$ and $j_1, \dots, j_k \in \mathbb{N}_0$ we have

$$\begin{aligned} & \mathbb{P}\{\lambda_{p_1}(\xi) \geq j_1, \dots, \lambda_{p_k}(\xi) \geq j_k\} = \mathbb{P}\{p_1^{j_1} \cdots p_k^{j_k} \text{ divides } \xi\} \\ &= \sum_{i=1}^{\infty} \mathbb{P}\{\xi = (p_1^{j_1} \cdots p_k^{j_k})i\} = (p_1^{j_1} \cdots p_k^{j_k})^{-\alpha} = p_1^{-\alpha j_1} \cdots p_k^{-\alpha j_k}. \end{aligned}$$

Thus, $(\lambda_p(\xi))_{p \in \mathcal{P}}$ are mutually independent and $\lambda_p(\xi)$ has a geometric distribution on \mathbb{N}_0 with parameter $p^{-\alpha}$, for every fixed $p \in \mathcal{P}$.

Example 2. For $\beta \in (0, 1)$, let $\mathbb{P}\{\xi = k\} = \beta^{k-1}(1 - \beta)$, $k \in \mathbb{N}$. Then

$$\mathbb{P}\{\lambda_p(\xi) \geq k\} = \frac{1 - \beta}{\beta} \sum_{j=1}^{\infty} \beta^{p^k j} = \frac{(1 - \beta)(\beta^{p^k - 1})}{1 - \beta p^k}, \quad k \in \mathbb{N}_0.$$

Example 3. Let $\text{Poi}(\lambda)$ be a random variable with the Poisson distribution with parameter λ and put

$$\mathbb{P}\{\xi = k\} = \mathbb{P}\{\text{Poi}(\lambda) = k | \text{Poi}(\lambda) \geq 1\} = (e^\lambda - 1)^{-1} \lambda^k / k!, \quad k \in \mathbb{N}.$$

Then

$$\begin{aligned} \mathbb{P}\{\lambda_p(\xi) \geq k\} &= (e^\lambda - 1)^{-1} \sum_{j=1}^{\infty} \lambda p^k j / (p^k j)! \\ &= \left({}_0F_{p^k} \left(; \frac{1}{p^k}, \frac{2}{p^k}, \dots, \frac{p^k - 1}{p^k}; \left(\frac{\lambda}{p^k} \right)^{p^k} \right) - 1 \right), \end{aligned} \tag{3}$$

where ${}_0F_{p^k}$ is the generalized hypergeometric function, see Chapter 16 in [10].

In all examples above, the distribution of $\lambda_p(\xi)$ for every fixed $p \in \mathcal{P}$ is extremely light-tailed. It is not that difficult to construct ‘weird’ distributions where all $\lambda_p(\xi)$ have infinite expectations.

Example 4. Let $(g_p)_{p \in \mathcal{P}}$ be any probability distribution supported by \mathcal{P} , $g_p > 0$, and $(t_k)_{k \in \mathbb{N}_0}$ any probability distribution on \mathbb{N} such that $\sum_{k=1}^{\infty} k t_k = \infty$ and $t_k > 0$. Define a probability distribution \mathfrak{h} on $\mathcal{Q} := \bigcup_{p \in \mathcal{P}} \{p, p^2, \dots\}$ by

$$\mathfrak{h}(\{p^k\}) = g_p t_k, \quad p \in \mathcal{P}, \quad k \in \mathbb{N}.$$

If ξ is a random variable with distribution \mathfrak{h} , then

$$\mathbb{P}\{\lambda_p(\xi) \geq k\} = g_p \sum_{j=k}^{\infty} t_j, \quad k \in \mathbb{N}, \quad p \in \mathcal{P},$$

which implies $\mathbb{E}[\lambda_p(\xi)] = g_p \sum_{k=1}^{\infty} k t_k = \infty$, $p \in \mathcal{P}$.

This example can be modified by taking $g := \sum_{p \in \mathcal{P}} g_p < 1$ and charging all points of $\mathbb{N} \setminus \mathcal{Q}$ (this set contains 1 and all integers having at least two different prime factors) with arbitrary positive masses of the total weight $1 - g$. The obtained probability distribution charges all points of \mathbb{N} and still possesses the property that all λ_p ’s have infinite expectations.

Let X be a random variable taking values in \mathbb{N} . Since

$$\log X = \sum_{p \in \mathcal{P}} \lambda_p(X) \log p,$$

we conclude that $\mathbb{E}[(\lambda_p(X))^k] < \infty$, for all $p \in \mathcal{P}$, whenever $\mathbb{E}[\log^k X] < \infty$, $k \in \mathbb{N}$. It is also clear that the converse implication is false in general. However, when $k = 1$ the inequality $\mathbb{E}[\log X] < \infty$ is in fact equivalent to $\sum_{p \in \mathcal{P}} \mathbb{E}[\lambda_p(X)] \log p < \infty$. As we have seen in the above examples, checking that $\mathbb{E}[(\lambda_p(X))^k] < \infty$ might be a much more difficult task than proving a stronger assumption $\mathbb{E}[\log^k X] < \infty$. Thus, we shall mostly work under moment conditions on $\log \xi$ and $\log \eta$.

Our standing assumption throughout the article is

$$\mu_\xi := \mathbb{E}[\log \xi] < \infty, \tag{4}$$

which, by the above reasoning, implies $\mathbb{E}[\lambda_p(\xi)] < \infty$, $p \in \mathcal{P}$.

2.2 Limit theorems for $S(p)$ and $T(p)$

From Donsker’s invariance principle we immediately obtain the following proposition. Let $D := D([0, \infty), \mathbb{R})$ be the Skorokhod space endowed with the standard J_1 -topology.

Proposition 1. *Assume that $\mathbb{E}[\log^2 \xi] \in (0, \infty)$. Then,*

$$\left(\left(\frac{S_{\lfloor ut \rfloor}(p) - ut \mathbb{E}[\lambda_p(\xi)]}{\sqrt{t}} \right)_{u \geq 0} \right)_{p \in \mathcal{P}} \implies ((W_p(u))_{u \geq 0})_{p \in \mathcal{P}}, \quad t \rightarrow \infty,$$

on the product space $D^{\mathbb{N}}$, where, for all $n \in \mathbb{N}$ and all $p_1 < p_2 < \dots < p_n$, $p_i \in \mathcal{P}$, $i \leq n$, $((W_{p_1}(u))_{u \geq 0}, \dots, (W_{p_n}(u))_{u \geq 0})$ is an n -dimensional centered Wiener process with covariance matrix $C = \|C_{i,j}\|_{1 \leq i,j \leq n}$ given by $C_{i,j} = C_{j,i} = \text{Cov}(\lambda_{p_i}(\xi), \lambda_{p_j}(\xi))$.

According to the proof of Proposition 1.3.13 in [6], see pp. 28–29 therein, the following holds true for the perturbed random walks $T(p)$, $p \in \mathcal{P}$.

Proposition 2. *Assume that $\mathbb{E}[\log^2 \xi] \in (0, \infty)$ and*

$$\lim_{t \rightarrow \infty} t^2 \mathbb{P}\{\lambda_p(\eta) \geq t\} = 0, \quad p \in \mathcal{P}. \tag{5}$$

Then,

$$\left(\left(\frac{T_{\lfloor ut \rfloor}(p) - ut \mathbb{E}[\lambda_p(\xi)]}{\sqrt{t}} \right)_{u \geq 0} \right)_{p \in \mathcal{P}} \implies ((W_p(u))_{u \geq 0})_{p \in \mathcal{P}}, \quad t \rightarrow \infty,$$

on the product space $D^{\mathbb{N}}$.

Remark 1. Since $\mathbb{P}\{\lambda_p(\eta) \log p \geq t\} \leq \mathbb{P}\{\log \eta \geq t\}$, the condition

$$\lim_{t \rightarrow \infty} t^2 \mathbb{P}\{\log \eta \geq t\} = 0 \tag{6}$$

is clearly sufficient for (5).

From the continuous mapping theorem under the assumptions of Proposition 2 we infer

$$\begin{aligned} & \left(\left(\frac{\max_{1 \leq k \leq \lfloor ut \rfloor} (T_k(p) - k \mathbb{E}[\lambda_p(\xi)])}{\sqrt{t}} \right)_{u \geq 0} \right)_{p \in \mathcal{P}} \\ & \implies \left(\left(\sup_{0 \leq v \leq u} W_p(v) \right)_{u \geq 0} \right)_{p \in \mathcal{P}}, \quad t \rightarrow \infty, \end{aligned} \tag{7}$$

see Proposition 1.3.13 in [6].

Formula (7), for a fixed $p \in \mathcal{P}$, belongs to the realm of limit theorems for the maximum of a single additive perturbed random walk. This circle of problems is well-understood, see Section 1.3.3 in [6] and [7], in the situation when the underlying additive standard random walk is *centered* and attracted to a stable Lévy process. In our setting the perturbed random walks $(T_k(p))_{k \in \mathbb{N}}$ and $(T_k(q))_{k \in \mathbb{N}}$ are dependent whenever $p, q \in \mathcal{P}$, $p \neq q$, which make derivation of the joint limit theorems harder and leads to various asymptotic regimes.

Note that (5) implies $\mathbb{E}[\lambda_p(\eta)] < \infty$ and (6) implies $\mathbb{E}[\log \eta] < \infty$. Theorem 5 below tells us that under such moment conditions and assuming also $\mathbb{E}[\log^2 \xi] < \infty$ the maxima $\max_{1 \leq k \leq n} T_k(p)$, $p \in \mathcal{P}$, of *noncentered* perturbed random walks $T(p)$ have the same behavior as $S_n(p)$, $p \in \mathcal{P}$ as $n \rightarrow \infty$.

Theorem 5. *Assume that $\mathbb{E}[\log^2 \xi] < \infty$ and $\mathbb{E}[\lambda_p(\eta)] < \infty$, $p \in \mathcal{P}$. Suppose further that*

$$\mathbb{P}\{\xi \text{ is divisible by } p\} = \mathbb{P}\{\lambda_p(\xi) > 0\} > 0, \quad p \in \mathcal{P}. \tag{8}$$

Then, as $t \rightarrow \infty$,

$$\left(\left(\frac{\max_{1 \leq k \leq [tu]} T_k(p) - \mathbb{E}[\lambda_p(\xi)]tu}{t^{1/2}} \right)_{u \geq 0} \right)_{p \in \mathcal{P}} \xrightarrow{\text{f.d.d.}} ((W_p(u))_{u \geq 0})_{p \in \mathcal{P}}. \tag{9}$$

Moreover, if also (5) holds for all $p \in \mathcal{P}$, then (9) holds on the product space $D^{\mathbb{N}}$.

Remark 2. If (8) holds only for some $\mathcal{P}_0 \subseteq \mathcal{P}$, then (9) holds with \mathcal{P}_0 instead of \mathcal{P} .

In the next result we shall assume that η dominates ξ in a sense that the asymptotic behavior of $\max_{1 \leq k \leq n} T_k(p)$ is regulated by the perturbations $(\lambda_p(\eta_k))_{k \leq n}$ for all $p \in \mathcal{P}_0$, where \mathcal{P}_0 is a finite subset of prime numbers and those p 's dominate all other primes.

Theorem 6. *Assume (4). Suppose further that there exists a finite set $\mathcal{P}_0 \subseteq \mathcal{P}$, $d := |\mathcal{P}_0|$, such that the distributional tail of $(\lambda_p(\eta))_{p \in \mathcal{P}_0}$ is regularly varying at infinity in the following sense. For some positive function $(a(t))_{t > 0}$ and a measure ν satisfying $\nu(\{x \in \mathbb{R}^d : \|x\| \geq r\}) = c \cdot r^{-\alpha}$, $c > 0$, $\alpha \in (0, 1)$, it holds*

$$t\mathbb{P}\{(a(t))^{-1}(\lambda_p(\eta))_{p \in \mathcal{P}_0} \in \cdot\} \xrightarrow{\nu} \nu(\cdot), \quad t \rightarrow \infty, \tag{10}$$

on the space of locally finite measures on $(0, \infty]^d$ endowed with the vague topology. Then

$$\left(\left(\frac{\max_{1 \leq k \leq [tu]} T_k(p)}{a(t)} \right)_{u \geq 0} \right)_{p \in \mathcal{P}_0} \xrightarrow{\text{f.d.d.}} ((M_p(u))_{u \geq 0})_{p \in \mathcal{P}_0}, \quad t \rightarrow \infty, \tag{11}$$

where $((M_p(u))_{u \geq 0})_{p \in \mathcal{P}_0}$ is a multivariate extreme process defined by

$$(M_p(u))_{p \in \mathcal{P}_0} = \sup_{k: t_k \leq u} y_k, \quad u \geq 0. \tag{12}$$

Here the pairs (t_k, y_k) are the atoms of a Poisson point process on $[0, \infty) \times (0, \infty]^d$ with the intensity measure $\mathbb{L} \otimes \nu$ and the supremum is taken coordinatewise. Moreover, suppose that $\mathbb{E}[\lambda_p(\eta)] < \infty$, for $p \in \mathcal{P} \setminus \mathcal{P}_0$. Then

$$\left(\left(\frac{\max_{1 \leq k \leq [tu]} T_k(p)}{a(t)} \right)_{u \geq 0} \right)_{p \in \mathcal{P} \setminus \mathcal{P}_0} \xrightarrow{\text{f.d.d.}} 0, \quad t \rightarrow \infty. \tag{13}$$

We shall deduce Theorems 5 and 6 in Section 3 by proving general limit results for coupled perturbed random walks.

2.3 Limit theorems for the LCM

The results from the previous section will be applied below to the analysis of

$$\mathfrak{P}_n := \text{LCM}(\{\Pi_1, \Pi_2, \dots, \Pi_n\}) \quad \text{and} \quad \mathfrak{T}_n := \text{LCM}(\{\Theta_1, \Theta_2, \dots, \Theta_n\}).$$

A moment's reflection shows that the analysis of \mathfrak{P}_n is trivial. Indeed, by definition, Π_{n-1} divides Π_n and thereupon $\mathfrak{P}_n = \Pi_n$ for $n \in \mathbb{N}$. Thus, assuming that $\sigma_\xi^2 := \text{Var}(\log \xi) \in (0, \infty)$, an application of the Donsker functional limit theorem yields

$$\left(\frac{\log \mathfrak{P}_{\lfloor tu \rfloor} - \mu_\xi t u}{t^{1/2}} \right)_{u \geq 0} \implies (\sigma_\xi W(u))_{u \geq 0}, \quad t \rightarrow \infty, \quad (14)$$

on the Skorokhod space D , where $(W(u))_{u \geq 0}$ is a standard Brownian motion and $\mu_\xi = \mathbb{E}[\log \xi]$ was defined in (4).

A simple structure of the sequence $(\mathfrak{P}_n)_{n \in \mathbb{N}}$ breaks down completely upon introducing the perturbations (η_k) , which makes the analysis of $(\mathfrak{T}_n)_{n \in \mathbb{N}}$ a much harder problem. As an illustration, consider the case $\xi = 1$ in which

$$\mathfrak{T}_n = \text{LCM}(\eta_1, \dots, \eta_n).$$

Thus, the problem encompasses, as a particular case, the investigation of the LCM of an independent sample. This itself constitutes a highly nontrivial challenge. Note that

$$\log \mathfrak{T}_n = \log \prod_{p \in \mathcal{P}} p^{\max_{1 \leq k \leq n} (\lambda_p(\xi_1) + \dots + \lambda_p(\xi_{k-1}) + \lambda_p(\eta_k))} = \sum_{p \in \mathcal{P}} \max_{1 \leq k \leq n} T_k(p) \log p,$$

which shows that the asymptotics of \mathfrak{T}_n is intimately connected with the behavior of $\max_{1 \leq k \leq n} T_k(p)$, $p \in \mathcal{P}$.

As one can guess from Theorem 5 in a ‘typical’ situation relation (14) holds with $\log \mathfrak{T}_{\lfloor tu \rfloor}$ replacing $\log \mathfrak{P}_{\lfloor tu \rfloor}$. The following heuristics suggest the right form of assumptions ensuring that perturbations $(\eta_k)_{k \in \mathbb{N}}$ have an asymptotically negligible impact on $\log \mathfrak{T}_n$. Take a prime $p \in \mathcal{P}$. Its contribution to $\log \mathfrak{T}_n$ (up to a factor $\log p$) is $\max_{1 \leq k \leq n} T_k(p)$. According to Theorem 5, this maximum is asymptotically the same as $S_n(p)$. However, as p gets large, the mean $\mathbb{E}[\lambda_p(\xi)]$ of the random walk $S_{n-1}(p)$ becomes small because of the identity

$$\sum_{p \in \mathcal{P}} \mathbb{E}[\lambda_p(\xi)] \log p = \mathbb{E}[\log \xi] < \infty.$$

Thus, for large $p \in \mathcal{P}$, the remainder $\max_{1 \leq k \leq n} T_k(p) - S_{n-1}(p)$ can, in principle, become larger than $S_{n-1}(p)$ itself if the tail of $\lambda_p(\eta)$ is sufficiently heavy. In order to rule out such a possibility, we introduce the deterministic sets

$$\mathcal{P}_1(n) := \{p \in \mathcal{P} : \mathbb{P}\{\lambda_p(\xi) > 0\} \geq n^{-1/2}\} \quad \text{and} \quad \mathcal{P}_2(n) := \mathcal{P} \setminus \mathcal{P}_1(n), \quad (15)$$

and bound the rate of growth of $\max_{1 \leq k \leq n} \lambda_p(\eta_k)$ for all $p \in \mathcal{P}_2(n)$. It is important to note that under the assumption (8) it holds

$$\min \mathcal{P}_2(n) = \min\{p \in \mathcal{P} : p \in \mathcal{P}_2(n)\}$$

$$= \min\{p \in \mathcal{P} : \mathbb{P}\{\lambda_p(\xi) > 0\} < n^{-1/2}\} \rightarrow \infty, \quad n \rightarrow \infty.$$

Therefore, if $\mathbb{E}[\log \xi] < \infty$ and (8) holds, then

$$\lim_{n \rightarrow \infty} \sum_{p \in \mathcal{P}_2(n)} \mathbb{E}[\lambda_p(\xi)] \log p = 0. \tag{16}$$

Theorem 7. Assume $\mathbb{E}[\log^2 \xi] < \infty$, $\mathbb{E}[\log \eta] < \infty$, (8) and the following two conditions:

$$\sum_{p \in \mathcal{P}} \mathbb{E}[\left((\lambda_p(\eta) - \lambda_p(\xi))^+\right)^2] \log p < \infty \tag{17}$$

and

$$\sum_{p \in \mathcal{P}_2(n)} \mathbb{E}[(\lambda_p(\eta) - \lambda_p(\xi))^+] \log p = o(n^{-1/2}), \quad n \rightarrow \infty. \tag{18}$$

Then

$$\left(\frac{\log \mathfrak{T}_{\lfloor tu \rfloor} - \mu_\xi tu}{t^{1/2}}\right)_{u \geq 0} \xrightarrow{\text{f.d.d.}} (\sigma_\xi W(u))_{u \geq 0}, \quad t \rightarrow \infty, \tag{19}$$

where $\mu_\xi = \mathbb{E}[\log \xi] < \infty$, $\sigma_\xi^2 = \text{Var}[\log \xi]$ and $(W(u))_{u \geq 0}$ is a standard Brownian motion.

Remark 3. If $\mathbb{E}[\log^2 \eta] < \infty$, then (17) holds true. Indeed, since we assume $\mathbb{E}[\log^2 \xi] < \infty$,

$$\begin{aligned} \mathbb{E}\left[\sum_{p \in \mathcal{P}} \left((\lambda_p(\eta) - \lambda_p(\xi))^+\right)^2 \log p\right] &\leq \mathbb{E}\left[\sum_{p \in \mathcal{P}} (\lambda_p^2(\eta) + \lambda_p^2(\xi)) \log p\right] \\ &\leq \frac{1}{\log 2} \mathbb{E}\left[\left(\sum_{p \in \mathcal{P}} \lambda_p(\eta) \log p\right)^2\right] + \mathbb{E}\left[\left(\sum_{p \in \mathcal{P}} \lambda_p(\xi) \log p\right)^2\right] \\ &= \frac{1}{\log 2} (\mathbb{E}[\log^2 \eta] + \mathbb{E}[\log^2 \xi]) < \infty. \end{aligned}$$

The condition (18) can be replaced by a stronger one which only involves the distribution of η , namely

$$\sum_{p \in \mathcal{P}_2(n)} \mathbb{E}[\lambda_p(\eta)] \log p = o(n^{-1/2}), \quad n \rightarrow \infty. \tag{20}$$

Taking into account (16) and the fact that $\mathbb{E}[\log \eta] < \infty$, the assumption (20) is nothing else but a condition of the speed of convergence of the series

$$\sum_{p \in \mathcal{P}} \mathbb{E}[\lambda_p(\eta)] \log p = \mathbb{E}[\log \eta].$$

Example 8. In the settings of Example 1, let ξ and η be arbitrarily dependent with

$$\mathbb{P}\{\xi = k\} = \frac{1}{\zeta(\alpha)k^\alpha}, \quad \mathbb{P}\{\eta = k\} = \frac{1}{\zeta(\beta)k^\beta}, \quad k \in \mathbb{N},$$

for some $\alpha, \beta > 1$. Note that $\mathbb{E}[\log^2 \xi] < \infty$ and $\mathbb{E}[\log^2 \eta] < \infty$. Direct calculations show that

$$\begin{aligned} \mathcal{P}_1(n) &= \{p \in \mathcal{P} : p^{-\alpha} \geq n^{-1/2}\} = \{p \in \mathcal{P} : p \leq n^{1/(2\alpha)}\}, \\ \mathcal{P}_2(n) &= \{p \in \mathcal{P} : p > n^{1/(2\alpha)}\}. \end{aligned}$$

From the chain of relations

$$\mathbb{E}[\lambda_p(\eta)] = \sum_{j \geq 1} \mathbb{P}\{\lambda_p(\eta) \geq j\} = \sum_{j \geq 1} p^{-\beta j} = \frac{p^{-\beta}}{1 - p^{-\beta}} \leq 2p^{-\beta},$$

and using the notation $\pi(x)$ for the number of primes smaller than x , we obtain

$$\begin{aligned} \sum_{p \in \mathcal{P}_2(n)} \mathbb{E}[\lambda_p(\eta)] \log p &\leq 2 \sum_{p \in \mathcal{P}, p > n^{1/(2\alpha)}} \frac{\log p}{p^\beta} = 2 \int_{(n^{1/(2\alpha)}, \infty)} \frac{\log x}{x^\beta} d\pi(x) \\ &\sim 2 \int_{n^{1/(2\alpha)}}^\infty \frac{\log x}{x^\beta} \frac{dx}{\log x} = \frac{2n^{(1-\beta)/(2\alpha)}}{\beta - 1}, \quad n \rightarrow \infty. \end{aligned}$$

Here the asymptotic equivalence follows from the prime number theorem and integration by parts, see, for example Eq. (16) in [3]. Thus, (20) holds if

$$\frac{1}{2} + \frac{1 - \beta}{2\alpha} < 0 \iff \alpha + 1 < \beta.$$

In the settings of Theorem 6 the situation is much simpler in a sense that almost no extra assumptions are needed to derive a limit theorem for \mathfrak{T}_n .

Theorem 9. *Under the same assumptions as in Theorem 6 and assuming additionally that*

$$\sum_{p \in \mathcal{P} \setminus \mathcal{P}_0} \mathbb{E}[\lambda_p(\eta)] \log p < \infty, \tag{21}$$

it holds

$$\left(\frac{\log \mathfrak{T}_{[tu]}}{a(t)}\right)_{u \geq 0} \xrightarrow{\text{f.d.d.}} \left(\sum_{p \in \mathcal{P}_0} M_p(u) \log p\right)_{u \geq 0}, \quad t \rightarrow \infty. \tag{22}$$

Note that in Theorem 9 it is allowed to take $\xi = 1$, which yields the following limit theorem for the LCM of an independent integer-valued random variables.

Corollary 1. *Under the same assumptions on η as in Theorem 6, it holds*

$$\left(\frac{\log \text{LCM}(\eta_1, \eta_2, \dots, \eta_{[tu]})}{a(t)}\right)_{u \geq 0} \xrightarrow{\text{f.d.d.}} \left(\sum_{p \in \mathcal{P}_0} M_p(u) \log p\right)_{u \geq 0}, \quad t \rightarrow \infty.$$

Remark 4. The results presented in Theorems 7 and 9 constitute a contribution to a popular topic in probabilistic number theory, namely, the asymptotic analysis of the LCM of various random sets. For random sets comprised of independent random variables uniformly distributed on $\{1, 2, \dots, n\}$ this problem has been addressed in [2–5, 9]. Some models with a more sophisticated dependence structure have been studied [1] and [8].

3 Limit theorems for coupled perturbed random walks

Theorems 5 and 6 will be derived from general limit theorems for the maxima of arbitrary additive perturbed random walks indexed by some parameters ranging in a countable set in the situation when the underlying additive standard random walks are positively divergent and attracted to a Brownian motion.

Let \mathcal{A} be a countable or finite set of real numbers and

$$\left((X_1(r), Y_1(r)) \right)_{r \in \mathcal{A}}, \quad \left((X_2(r), Y_2(r)) \right)_{r \in \mathcal{A}}, \dots$$

be independent copies of an $\mathbb{R}^{2 \times |\mathcal{A}|}$ random vector $(X(r), Y(r))_{r \in \mathcal{A}}$ with arbitrarily dependent components. For each $r \in \mathcal{A}$, the sequence $(S_k^*(r))_{k \in \mathbb{N}_0}$ given by

$$S_0^*(r) := 0, \quad S_k^*(r) := X_1(r) + \dots + X_k(r), \quad k \in \mathbb{N},$$

is an additive standard random walk. For each $r \in \mathcal{A}$, the sequence $(T_k^*(r))_{k \in \mathbb{N}}$ defined by

$$T_k^*(r) := S_{k-1}^*(r) + Y_k(r), \quad k \in \mathbb{N},$$

is an additive perturbed random walk. The sequence $((T_k^*(r))_{k \in \mathbb{N}})_{r \in \mathcal{A}}$ is a collection of (generally) dependent additive perturbed random walks.

Proposition 3. *Assume that, for each $r \in \mathcal{A}$, $\mu(r) := \mathbb{E}[X(r)] \in (0, \infty)$, $\text{Var}[X(r)] \in [0, \infty)$ and $\mathbb{E}[Y(r)] < \infty$. Then*

$$\left(\left(\frac{\max_{1 \leq k \leq \lfloor tu \rfloor} T_k^*(r) - \mu(r)tu}{t^{1/2}} \right)_{u \geq 0} \right)_{r \in \mathcal{A}} \xrightarrow{\text{f.d.d.}} \left((W_r(u))_{u \geq 0} \right)_{r \in \mathcal{A}}, \quad t \rightarrow \infty, \tag{23}$$

where, for all $n \in \mathbb{N}$ and arbitrary $r_1 < r_2 < \dots < r_n$ with $r_i \in \mathcal{A}$, $i \leq n$, $((W_{r_1}(u))_{u \geq 0}, \dots, (W_{r_n}(u))_{u \geq 0})$ is an n -dimensional centered Wiener process with covariance matrix $C = \|C_{i,j}\|_{1 \leq i,j \leq n}$ with the entries $C_{i,j} = C_{j,i} = \text{Cov}(X(r_i), X(r_j))$.

Proof. We shall prove an equivalent statement that, as $t \rightarrow \infty$,

$$\left(\left(\frac{\max_{0 \leq k \leq \lfloor tu \rfloor} T_{k+1}^*(r) - \mu(r)tu}{t^{1/2}} \right)_{u \geq 0} \right)_{r \in \mathcal{A}} \xrightarrow{\text{f.d.d.}} \left((W_r(u))_{u \geq 0} \right)_{r \in \mathcal{A}},$$

which differs from (23) by a shift of the subscript k . By the multidimensional Donsker theorem,

$$\left(\left(\frac{S_{\lfloor tu \rfloor}^*(r) - \mu(r)tu}{t^{1/2}} \right)_{u \geq 0} \right)_{r \in \mathcal{A}} \implies \left((W_r(u))_{u \geq 0} \right)_{r \in \mathcal{A}}, \quad t \rightarrow \infty, \tag{24}$$

in the product topology of $D^{\mathbb{N}}$. Fix any $r \in \mathcal{A}$ and write

$$\begin{aligned} & \max_{0 \leq k \leq \lfloor tu \rfloor} T_{k+1}^*(r) - \mu(r)tu \\ &= \max_{0 \leq k \leq \lfloor tu \rfloor} (S_k^*(r) - S_{\lfloor tu \rfloor}^*(r) + Y_{k+1}(r)) + S_{\lfloor tu \rfloor}^*(r) - \mu(r)tu. \end{aligned} \tag{25}$$

In view of (24) the proof is complete once we can show that

$$n^{-1/2} \left(\max_{0 \leq k \leq n} (S_k^*(r) - S_n^*(r) + Y_{k+1}(r)) \right) \xrightarrow{\mathbb{P}} 0, \quad n \rightarrow \infty. \quad (26)$$

Let $(X_0(r), Y_0(r))$ be a copy of $(X(r), Y(r))$ which is independent of the vector $(X_k(r), Y_k(r))_{k \in \mathbb{N}}$. Since the collection

$$((X_1(r), Y_1(r)), \dots, (X_{n+1}(r), Y_{n+1}(r)))$$

has the same distribution as

$$((X_n(r), Y_n(r)), \dots, (X_0(r), Y_0(r))),$$

the variable

$$\max_{0 \leq k \leq n} (S_k^*(r) - S_n^*(r) + Y_{k+1}(r))$$

has the same distribution as

$$\max \left(Y_0(r), \max_{0 \leq k \leq n-1} (-S_k^*(r) + Y_{k+1}(r) - X_{k+1}(r)) \right).$$

By assumption, $\mathbb{E}(-S_1^*(r)) \in (-\infty, 0)$ and $\mathbb{E}(Y(r) - X(r))^+ < \infty$. Hence, by Theorem 1.2.1 and Remark 1.2.3 in [6],

$$\lim_{k \rightarrow \infty} (-S_k^*(r) + Y_{k+1}(r) - X_{k+1}(r)) = -\infty \quad \text{a.s.}$$

As a consequence, the a.s. limit

$$\begin{aligned} & \lim_{n \rightarrow \infty} \max(Y_0(r), \max_{0 \leq k \leq n-1} (-S_k^*(r) + Y_{k+1}(r) - X_{k+1}(r))) \\ &= \max(Y_0(r), \max_{k \geq 0} (-S_k^*(r) + Y_{k+1}(r) - X_{k+1}(r))) \end{aligned}$$

is a.s. finite. This completes the proof of (26). □

Remark 5. Proposition 3 tells us that fluctuations of $\max_{1 \leq k \leq \lfloor tu \rfloor} T_k^*(r)$ on the level of finite-dimensional distributions are driven by the Brownian fluctuations of $S_{\lfloor tu \rfloor}^*(r)$. According to formula (25), a functional version of this statement would be true if we could check that, for every fixed $T > 0$,

$$t^{-1/2} \sup_{u \in [0, T]} \max_{0 \leq k \leq \lfloor tu \rfloor} (S_k^*(r) - S_{\lfloor tu \rfloor}^*(r) + Y_{k+1}(r)) \xrightarrow{\mathbb{P}} 0, \quad t \rightarrow \infty.$$

But the left-hand side is bounded from below by

$$t^{-1/2} \sup_{u \in [0, T]} Y_{\lfloor tu \rfloor + 1}(r) = t^{-1/2} \max_{0 \leq k \leq \lfloor Tt \rfloor + 1} Y_k(r).$$

Under the sole assumption $\mathbb{E}[Y(r)] < \infty$ this maximum does not converge to zero in probability, as $t \rightarrow \infty$. Thus, under the standing assumptions of Proposition 3 the functional convergence does not hold.

Proof of Theorem 5. To deduce the finite-dimensional convergence (9) we apply Proposition 3 with $\mathcal{A} = \mathcal{P}$, $X(p) = \lambda_p(\xi)$ and $Y(p) = \lambda_p(\eta)$. The assumption (8) in conjunction with $\mathbb{E}[\log^2 \xi] < \infty$ implies that $\mathbb{E}[\lambda_p(\xi)] \in (0, \infty)$ and $\text{Var}[\lambda_p(\xi)] \in [0, \infty)$, for all $p \in \mathcal{P}$.

Suppose that (5) holds true for all $p \in \mathcal{P}$. Fix $p \in \mathcal{P}$, $t > 0$, and note that by the subadditivity of the supremum and the fact that $(S_k(p))_{k \in \mathbb{N}_0}$ is nondecreasing we have

$$S_{\lfloor tu \rfloor - 1}(p) \leq \max_{1 \leq k \leq \lfloor tu \rfloor} T_k(p) \leq S_{\lfloor tu \rfloor - 1}(p) + \max_{1 \leq k \leq \lfloor tu \rfloor} \lambda_p(\eta_k), \quad u \geq 0. \quad (27)$$

Assumption (5) implies that, for every fixed $T > 0$,

$$t^{-1/2} \sup_{u \in [0, T]} \max_{1 \leq k \leq \lfloor tu \rfloor} \lambda_p(\eta_k) = t^{-1/2} \max_{1 \leq k \leq \lfloor tT \rfloor} \lambda_p(\eta_k) \xrightarrow{\mathbb{P}} 0, \quad t \rightarrow \infty.$$

By Proposition 1 and taking into account (27) this means that (9) holds true on the product space $D^{\mathbb{N}}$. \square

Proposition 4. Assume $\mathbb{E}[X(r)] < \infty$, $r \in \mathcal{A}$. Assume further that there exists a finite set $\mathcal{A}_0 \subseteq \mathcal{A}$, $d := |\mathcal{A}_0|$, such that the distributional tail of $(Y(r))_{r \in \mathcal{A}_0}$ is regularly varying at infinity in the following sense. For some positive function $(a(t))_{t > 0}$ and a measure ν satisfying $\nu(\{x \in \mathbb{R}^d : \|x\| \geq r\}) = c \cdot r^{-\alpha}$, $c > 0$, $\alpha \in (0, 1)$, it holds

$$t\mathbb{P}\{(a(t))^{-1}(Y(r))_{r \in \mathcal{A}_0} \in \cdot\} \xrightarrow{\nu} \nu(\cdot), \quad t \rightarrow \infty, \quad (28)$$

on the space of locally finite measures on $(0, \infty]^d$ endowed with the vague topology. Then

$$\left(\left(\frac{\max_{1 \leq k \leq \lfloor tu \rfloor} T_k^*(r)}{a(t)} \right)_{u \geq 0} \right)_{r \in \mathcal{A}_0} \xrightarrow{\text{f.d.d.}} ((M_r(u))_{u \geq 0})_{r \in \mathcal{A}_0}, \quad t \rightarrow \infty, \quad (29)$$

where $((M_r(u))_{u \geq 0})_{r \in \mathcal{A}_0}$ is defined as in (12). If $\mathbb{E}[|Y(r)|] < \infty$, for $r \in \mathcal{A} \setminus \mathcal{A}_0$, then also

$$\left(\left(\frac{\max_{1 \leq k \leq \lfloor tu \rfloor} T_k^*(r)}{a(t)} \right)_{u \geq 0} \right)_{r \in \mathcal{A} \setminus \mathcal{A}_0} \xrightarrow{\text{f.d.d.}} 0, \quad t \rightarrow \infty. \quad (30)$$

Proof. According to Corollary 5.18 in [11]

$$\left(\left(\frac{\max_{1 \leq k \leq \lfloor tu \rfloor} Y_k(r)}{a(t)} \right)_{u \geq 0} \right)_{r \in \mathcal{A}_0} \implies ((M_r(u))_{u \geq 0})_{r \in \mathcal{A}_0}, \quad t \rightarrow \infty,$$

in the product topology of $D^{\mathbb{N}}$. The function $(a(t))_{t \geq 0}$ is regularly varying at infinity with index $1/\alpha > 1$. Thus, by the law of large numbers, for all $r \in \mathcal{A}$,

$$\left(\frac{\min_{1 \leq k \leq \lfloor tu \rfloor} S_{k-1}^*(r)}{a(t)} \right)_{u \geq 0} \xrightarrow{\text{f.d.d.}} 0, \quad t \rightarrow \infty, \quad (31)$$

$$\left(\frac{\max_{1 \leq k \leq \lfloor tu \rfloor} S_{k-1}^*(r)}{a(t)} \right)_{u \geq 0} \xrightarrow{\text{f.d.d.}} 0, \quad t \rightarrow \infty, \quad (32)$$

and (29) follows from the inequalities

$$\begin{aligned} \min_{1 \leq k \leq \lfloor tu \rfloor} S_{k-1}^*(r) + \max_{1 \leq k \leq \lfloor tu \rfloor} Y_k(r) &\leq \max_{1 \leq k \leq \lfloor tu \rfloor} T_k^*(r) \\ &\leq \max_{1 \leq k \leq \lfloor tu \rfloor} S_{k-1}^*(r) + \max_{1 \leq k \leq \lfloor tu \rfloor} Y_k(r). \end{aligned}$$

In view of (31) and (32), to prove (30) it suffices to check that

$$\left(\left(\frac{\max_{1 \leq k \leq \lfloor tu \rfloor} Y_k(r)}{a(t)} \right)_{u \geq 0} \right) \xrightarrow{\text{f.d.d.}} 0, \quad t \rightarrow \infty,$$

for every fixed $r \in \mathcal{A} \setminus \mathcal{A}_0$. This, in turn, follows from

$$\frac{Y_n(r)}{n} \xrightarrow{\text{a.s.}} 0, \quad n \rightarrow \infty, \quad r \in \mathcal{A} \setminus \mathcal{A}_0,$$

which is a consequence of the assumption $\mathbb{E}[|Y(r)|] < \infty$, $r \in \mathcal{A} \setminus \mathcal{A}_0$, and the Borel–Cantelli lemma. □

Proof of Theorem 6. Follows immediately from Proposition 4 applied with $\mathcal{A} = \mathcal{P}$, $X(p) = \lambda_p(\xi)$ and $Y(p) = \lambda_p(\eta)$. □

4 Proof of Theorem 7

We aim at proving that

$$\frac{\sum_{p \in \mathcal{P}} (\max_{1 \leq k \leq n} T_k(p) - S_{n-1}(p)) \log p}{\sqrt{n}} \xrightarrow{\mathbb{P}} 0, \quad n \rightarrow \infty, \quad (33)$$

which together with the relation

$$\sum_{p \in \mathcal{P}} S_n(p) \log p = \log \Pi_n = \log \mathfrak{P}_n, \quad n \in \mathbb{N},$$

implies Theorem 7 by the Slutsky lemma and (14).

Let (ξ_0, η_0) be an independent copy of (ξ, η) which is also independent of $(\xi_n, \eta_n)_{n \in \mathbb{N}}$. By the same reasoning as we have used in the proof of (26) we obtain

$$\begin{aligned} &\left(\max_{1 \leq k \leq n} T_k(p) - S_{n-1}(p) \right)_{p \in \mathcal{P}} \\ &\stackrel{d}{=} \left(\max \left(\lambda_p(\eta_0), \max_{1 \leq k < n} (\lambda_p(\eta_k) - \lambda_p(\xi_k) - S_{k-1}(p)) \right) \right)_{p \in \mathcal{P}}. \end{aligned} \quad (34)$$

Taking into account

$$\sum_{p \in \mathcal{P}} \lambda_p(\eta_0) \log p = \log \eta_0,$$

we see that (33) is a consequence of

$$\frac{\sum_{p \in \mathcal{P}} \max_{1 \leq k < n} (\lambda_p(\eta_k) - \lambda_p(\xi_k) - S_{k-1}(p))^+ \log p}{\sqrt{n}} \xrightarrow{\mathbb{P}} 0, \quad n \rightarrow \infty. \quad (35)$$

Since, for every fixed $p \in \mathcal{P}$,

$$\max_{k \geq 1} (\lambda_p(\eta_k) - \lambda_p(\xi_k) - S_{k-1}(p))^+ < \infty \quad \text{a.s.} \tag{36}$$

by assumption (8), it suffices to check that, for every fixed $\varepsilon > 0$,

$$\lim_{M \rightarrow \infty} \limsup_{n \rightarrow \infty} \mathbb{P} \left\{ \sum_{p \in \mathcal{P}, p > M} \max_{1 \leq k < n} (\lambda_p(\eta_k) - \lambda_p(\xi_k) - S_{k-1}(p))^+ \log p > \varepsilon \sqrt{n} \right\}. \tag{37}$$

In order to check (37), we divide the sum into two disjoint parts with summations over $\mathcal{P}_1(n)$ and $\mathcal{P}_2(n)$. For the first sum, by Markov’s inequality, we obtain

$$\begin{aligned} & \mathbb{P} \left\{ \sum_{p \in \mathcal{P}_1(n), p > M} \max_{1 \leq k < n} (\lambda_p(\eta_k) - \lambda_p(\xi_k) - S_{k-1}(p))^+ \log p > \varepsilon \sqrt{n} / 2 \right\} \\ & \leq \frac{2}{\varepsilon \sqrt{n}} \sum_{p \in \mathcal{P}_1(n), p > M} \mathbb{E} \left(\max_{1 \leq k < n} (\lambda_p(\eta_k) - \lambda_p(\xi_k) - S_{k-1}(p))^+ \right) \log p \\ & \leq \frac{2}{\varepsilon \sqrt{n}} \sum_{p \in \mathcal{P}_1(n), p > M} \log p \sum_{k \geq 1} \mathbb{E} (\lambda_p(\eta_k) - \lambda_p(\xi_k) - S_{k-1}(p))^+ \\ & = \frac{2}{\varepsilon \sqrt{n}} \sum_{p \in \mathcal{P}_1(n), p > M} \log p \sum_{j \geq 1} \mathbb{P} \{ \lambda_p(\eta) - \lambda_p(\xi) = j \} \sum_{k \geq 1} \mathbb{E} (j - S_{k-1}(p))^+ \\ & \leq \frac{2}{\varepsilon \sqrt{n}} \sum_{p \in \mathcal{P}_1(n), p > M} \log p \sum_{j \geq 1} j \mathbb{P} \{ \lambda_p(\eta) - \lambda_p(\xi) = j \} \sum_{k \geq 0} \mathbb{P} \{ S_k(p) \leq j \} \\ & \leq \frac{2}{\varepsilon \sqrt{n}} \sum_{p \in \mathcal{P}_1(n), p > M} \log p \sum_{j \geq 1} j \mathbb{P} \{ \lambda_p(\eta) - \lambda_p(\xi) = j \} \frac{2j}{\mathbb{E}[(\lambda_p(\xi) \wedge j)]}, \end{aligned}$$

where the last estimate is a consequence of Erickson’s inequality for renewal functions, see Eq. (6.5) in [6]. Further, since for $p \in \mathcal{P}_1(n)$,

$$\mathbb{E}[(\lambda_p(\xi) \wedge j)] \geq \mathbb{P} \{ \lambda_p(\xi) \geq 1 \} = \mathbb{P} \{ \lambda_p(\xi) > 0 \} \geq n^{-1/2},$$

we obtain

$$\begin{aligned} & \mathbb{P} \left\{ \sum_{p \in \mathcal{P}_1(n), p > M} \max_{1 \leq k < n} (\lambda_p(\eta_k) - \lambda_p(\xi_k) - S_{k-1}(p))^+ \log p > \varepsilon \sqrt{n} / 2 \right\} \\ & \leq \frac{4}{\varepsilon} \sum_{p \in \mathcal{P}_1(n), p > M} \log p \mathbb{E} [((\lambda_p(\eta) - \lambda_p(\xi))^+)^2] \\ & \leq \frac{4}{\varepsilon} \sum_{p \in \mathcal{P}, p > M} \log p \mathbb{E} [((\lambda_p(\eta) - \lambda_p(\xi))^+)^2]. \end{aligned}$$

The right-hand side converges to 0, as $M \rightarrow \infty$ by (17). For the sum over $\mathcal{P}_2(n)$ the derivation is simpler. By Markov’s inequality

$$\mathbb{P} \left\{ \sum_{p \in \mathcal{P}_2(n), p > M} \max_{1 \leq k < n} (\lambda_p(\eta_k) - \lambda_p(\xi_k) - S_{k-1}(p))^+ \log p > \varepsilon \sqrt{n} / 2 \right\}$$

$$\begin{aligned} &\leq \frac{2}{\varepsilon\sqrt{n}} \mathbb{E} \left[\sum_{p \in \mathcal{P}_2(n), p > M} \max_{1 \leq k < n} (\lambda_p(\eta_k) - \lambda_p(\xi_k) - S_{k-1}(p))^+ \log p \right] \\ &\leq \frac{2n}{\varepsilon\sqrt{n}} \mathbb{E} \left[\sum_{p \in \mathcal{P}_2(n), p > M} (\lambda_p(\eta_k) - \lambda_p(\xi_k))^+ \log p \right], \end{aligned}$$

and the right-hand side tends to zero as $n \rightarrow \infty$ in view of (18). The proof is complete.

5 Proof of Theorem 9

From Theorem 6 with the aid of the continuous mapping theorem we conclude that

$$\left(\frac{\sum_{p \in \mathcal{P}_0} \max_{1 \leq k \leq \lfloor tu \rfloor} T_k(p) \log p}{a(t)} \right)_{u \geq 0} \xrightarrow{\text{f.d.d.}} \left(\sum_{p \in \mathcal{P}_0} M_p(u) \log p \right)_{u \geq 0},$$

as $t \rightarrow \infty$. It suffices to check

$$\left(\frac{\sum_{p \in \mathcal{P} \setminus \mathcal{P}_0} \max_{1 \leq k \leq \lfloor tu \rfloor} T_k(p) \log p}{a(t)} \right)_{u \geq 0} \xrightarrow{\text{f.d.d.}} 0, \quad t \rightarrow \infty. \tag{38}$$

Since $(a(t))$ is regularly varying at infinity, (38) follows from

$$\frac{\sum_{p \in \mathcal{P} \setminus \mathcal{P}_0} \mathbb{E}[\max_{1 \leq k \leq n} T_k(p)] \log p}{a(n)} \rightarrow 0, \quad n \rightarrow \infty, \tag{39}$$

by Markov’s inequality. To check the latter, note that

$$\begin{aligned} \sum_{p \in \mathcal{P} \setminus \mathcal{P}_0} \mathbb{E} \left[\max_{1 \leq k \leq n} T_k(p) \right] \log p &\leq \sum_{p \in \mathcal{P} \setminus \mathcal{P}_0} \mathbb{E} \left[S_{n-1}(p) + \max_{1 \leq k \leq n} \lambda_p(\eta_k) \right] \log p \\ &\leq (n-1) \sum_{p \in \mathcal{P} \setminus \mathcal{P}_0} \mathbb{E}[\lambda_p(\xi)] \log p + n \sum_{p \in \mathcal{P} \setminus \mathcal{P}_0} \mathbb{E}[\lambda_p(\eta)] \log p \\ &\leq (n-1) \mathbb{E}[\log \xi] + n \sum_{p \in \mathcal{P} \setminus \mathcal{P}_0} \mathbb{E}[\lambda_p(\eta)] \log p = O(n), \quad n \rightarrow \infty, \end{aligned}$$

where we have used the inequality $\mathbb{E}[\log \xi] < \infty$ and the assumption (21). Using that $\alpha \in (0, 1)$ and $(a(t))$ is regularly varying at infinity with index $1/\alpha$, we obtain (39).

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